

# Real-Time Trajectory Generation Algorithms for Uncertain Dynamical Systems Using Covariance Steering

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# Motivation: High-Precision Control of Uncertain Dynamical Systems



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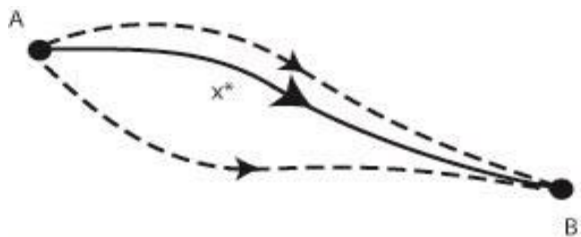


Image taken from Raytheon.com

- Modern/future control systems are required to perform control tasks with high precision under extremely uncertain conditions
- **Main challenge:** Accurate models for system dynamics and / or system-environment interactions may not be available a priori
- **Goal:** Design control algorithms for high-precision control of uncertain systems based on the idea of directly controlling the effects of model uncertainty and unknown disturbances on the control system

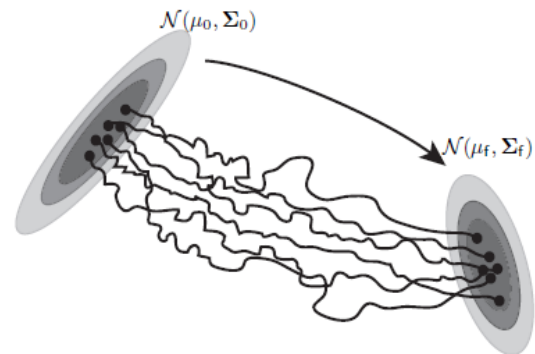
# Motivation: A New Perspective on Trajectory Generation / Optimization Problems

- **Deterministic Trajectory Generation/Optimization:** Steer a system “from point A to point B” while minimizing a relevant performance index without violating certain input and / or state constraints
- **Trajectory Optimization for Uncertain Systems:**
  - system model is stochastic and may not be available a priori
  - boundary conditions should be *probabilistic* rather than deterministic
- **Distribution / Covariance Steering:** Special class of stochastic trajectory optimization problems in which the goal is to drive the state mean and state covariance to prescribed quantities



Standard (deterministic) trajectory generation problem

v/s



Covariance Steering Problem

# **PART I: Model-Based Nonlinear Covariance Steering**



# Covariance Steering for Discrete-Time Stochastic Nonlinear Systems

- Given the discrete-time stochastic nonlinear system:

$$x(t + 1) = f(x(t), u(t)) + w(t), \quad x(0) = x_0 \quad (1)$$

- $\mathbb{E}[x_0] = \mu_0$ ,  $\text{cov}(x_0, x_0) = \Sigma_0$ , where  $\mu_0 \in \mathbb{R}^n$ ,  $\Sigma_0 = \Sigma_0' > 0$  (given)
- $\{x(t): t \in [0, N]_d\}$ : state process,  $\{u(t): t \in [0, N - 1]_d\}$ : input process,
- $\{w(t): t \in [0, N - 1]_d\}$ : sequence of i.i.d. (normal) random variables

$$\mathbb{E}(w(t)) = 0, \mathbb{E}(w(t)w(\tau)') = \delta(t, \tau)\mathbf{W}, \text{ where } \mathbf{W} = \mathbf{W}' \geq 0$$

- Nonlinear Covariance Steering Problem:** Given  $\mu_d \in \mathbb{R}^n$ ,  $\Sigma_d = \Sigma_d' > 0$ , find a control policy

$$\boldsymbol{\omega} := \{\varphi_0(x), \varphi_1(x), \dots, \varphi_{N-1}(x)\}$$

that will steer the system (1) to a state  $x(N) = x_f$  with

$$\mathbb{E}[x_f] = \mu_d, \quad \text{cov}(x_f, x_f) = \Sigma_d \quad (2)$$

while minimizing a relevant performance index:  $J(u) = \mathbb{E}[\sum_{t=0}^{N-1} |u(t)|^2]$



## Successive Linearization of Dynamics

- At stage  $t = k$  the nonlinear system is linearized around  $(\mu_k, \nu_k)$ , where  $\mu_k$  and  $\nu_k$  are (approximations of) the state mean and input mean at  $t = k$

$$z(t + 1) = \mathbf{A}_k(z(t) - \mu_k) + \mathbf{B}_k(u(t) - \nu_k) + r_k + w(t),$$

$$\mathbf{A}_k = f_x(\mu_k, \nu_k), \quad \mathbf{B}_k = f_u(\mu_k, \nu_k), \quad r_k = f(\mu_k, \nu_k)$$

- Equivalently,

$$z(t + 1) = \mathbf{A}_k z(t) + \mathbf{B}_k u(t) + d_k + w(t), \quad t \in [k, N - 1]_d,$$

$$\text{where } d_k = -\mathbf{A}_k \mu_k - \mathbf{B}_k \nu_k + r_k$$

- Boundary conditions for covariance steering:**  $z(k) = z_k$  where  $\mathbb{E}[z_k] = \mu_k$ ,  $\text{COV}(z_k, z_k) = \Sigma_k$
- Remark:** The  $k$ -th state space model of the successive linearization approach is computed based on information available at stage  $t = k$



# Alternative Linearization based on a Reference Trajectory

- Alternatively, linearize around reference state and input trajectories \*  $\{\bar{z}(t): t \in [0, N]_d\}$  and  $\{\bar{u}(t): t \in [0, N - 1]_d\}$ , respectively:

$$z(t + 1) = \mathbf{A}(t)(z(t) - \bar{z}(t)) + \mathbf{B}(t)(u(t) - \bar{u}(t)) + r_k + w(t),$$

$$\mathbf{A}(t) = f_x(\bar{z}(t), \bar{u}(t)),$$

$$\mathbf{B}(t) = f_u(\bar{z}(t), \bar{u}(t)),$$

$$r(t) = f(\bar{z}(t), \bar{u}(t)) - \mathbf{A}(t)\bar{z}(t) - \mathbf{B}(t)\bar{u}(t)$$

- The above linearization corresponds to a *single* time-varying state space model and is based on information available at  $t = 0$  (computed off-line)
- By contrast, the successive linearization approach relies on the successive computation of time-invariant state space models (computed on-the-fly, i.e., a new time-invariant model at a time)

Ridderhof, J., Okamoto, K. and Tsiotras, P., 2019, December. Nonlinear uncertainty control with iterative covariance steering. In 2019 IEEE 58th Conference on Decision and Control (CDC) (pp. 3484-3490). IEEE

# Linear Covariance Steering Problem

**Problem:** Find a control policy  $\pi_k = \{\phi_k(t, z): t \in [k, N - 1]\}$ , where

$$\phi_k(t, z) := v_k(t) + \mathbf{K}_k(t)z,$$

that minimizes the performance index:

$$J(\pi_k) := \mathbb{E} \left[ \sum_{t=k}^{N-1} |\phi_k(t, z)|^2 \right]$$

subject to the dynamic constraints:

$$z(t + 1) = \mathbf{A}_k z(t) + \mathbf{B}_k u(t) + d_k + w(t), \quad t \in [k, N - 1]_d$$

and the boundary constraints:

$$\mathbb{E}[z(k)] = \mu_k \quad \text{cov}(z(k), z(k)) = \mathbf{\Sigma}_k$$

$$\mathbb{E}[z(N)] = \mu_d \quad \text{cov}(z(N), z(N)) = \mathbf{\Sigma}_d$$

for given  $\mu_k, \mu_d \in \mathbb{R}^n$  and  $\mathbf{\Sigma}_k, \mathbf{\Sigma}_d \in \mathbb{S}_n^{++}$  (positive definite matrices)

**Remark 1:** Subscript  $k$  indicates dependence on information available at stage  $t = k$  (i.e., the policy  $\pi_k$  consists of control laws for  $t \in [k, N - 1]$  based on information available at stage  $t = k$ )





# Solution to Linear Covariance Steering Problem (SDP Convex Optimization Formulation)

- The  $k$ -th linearized covariance steering can be reduced (relaxed) to a semi-definite program (SDP)\*:

$$\Sigma_d - \text{cov}(z(N), z(N)) \succeq \mathbf{0} \quad (\text{LMI constraint})$$

- The solution to the latter problem will furnish the sequence of gains  $\mathcal{K}_k = \{\mathbf{K}_k(t) : t \in [k, N - 1]_d\}$  and  $\mathbf{v}_k = \{v_k(t) : t \in [k, N - 1]_d\}$  that will solve the  $k$ -th linearized covariance steering problem

$$\pi_k = \pi_k(\mathcal{K}_k, \mathbf{v}_k) = \{\phi_k(k, z), \dots, \phi_k(N - 1, z)\},$$

- $\phi_k(t, z) := v_k(t) + \mathbf{K}_k(t)z$  (state feedback)
- $\phi_k(t, Z) := v_k(t) + \sum_{\tau=k}^t \mathbf{K}_k(\tau, t)z(\tau)$  where  $Z = \{z(k), \dots, z(t)\}$   
(history-based state-feedback)

\* Bakolas, E. "Finite-Horizon Covariance Control for Discrete-Time Stochastic Linear Systems Subject to Input Constraints," *Automatica*, vol. 91, no. 5, pp. 61-68, 2018.



# Solution to the Linear Covariance Steering Problem (DCP Formulation)

- The computationally tractable SDP formulation of the covariance steering problem relies on the assumption that the latter problem admits a solution for the given (hard) terminal conditions on mean and covariance
- An alternative formulation is based on replacing the hard constraints:

$$\text{cov}(z(N), z(N)) = \Sigma_d \quad \mathbb{E}[z(N)] = \mu_d$$

with a terminal cost term measuring the distance between  $\mathcal{N}(\mu_d, \Sigma_d)$  and the Gaussian approximation  $\mathcal{N}(\mu_N, \Sigma_N)$  where  $\mu_N = \mathbb{E}[z(N)]$ , and  $\Sigma_N = \text{cov}(z(N), z(N))$  of the terminal state distribution:

$$J(\pi_k) := \gamma \mathcal{W}^2(\rho_d, \rho_N) + \mathbb{E} \left[ \sum_{t=k}^{N-1} |\phi_k(t, z)|^2 \right]$$

where  $\gamma > 0$  and  $\mathcal{W}(\rho_d, \rho_N)$  is the Wasserstein distance between the desired and the actual terminal state distributions\*

\*Halder, Abhishek, and Eric DB Wendel. "Finite horizon linear quadratic Gaussian density regulator with Wasserstein terminal cost." In 2016 American Control Conference (ACC), pp. 7249-7254. IEEE, 2016.

## Solution to $k$ -th Linearized Covariance Steering Problem (Difference of Convex Functions Program Formulation)

- Let the actual terminal state and desired distributions be approximated by Gaussian distributions with densities  $\rho_N, \rho_d$ , respectively. Then

$$\mathcal{W}^2(\rho_d, \rho_N) = |\mu_d - \mu_N|^2 + \text{tr} \left( \Sigma_d + \Sigma_N - 2 \left( \Sigma_N^{1/2} \Sigma_d \Sigma_N^{1/2} \right)^{1/2} \right)$$

- It can be shown that  $J(\pi_k)$  can be expressed as a difference of two convex functions\*
- The covariance steering problem with terminal cost turns out to be a *Difference of Convex functions Program* (DCP)
- DCPs are computationally tractable (one can solve them by using, for instance, the so-called convex / concave procedure)

\* I. Balci and E. Bakolas, "Covariance Steering of Discrete-Time Stochastic Linear Systems Based on Wasserstein Distance Terminal Cost," in *IEEE Control Systems Letters*, doi: 10.1109/LCSYS.2020.3047132

# Unscented Transform for One-Stage Prediction of State Mean and Covariance

- Given a feedback control policy  $\pi_k$ , the closed-loop dynamics of the original (nonlinear system) are described by the following equation:

$$x(t + 1) = f_{cl}^k(t, x(t)) + w(t), \quad t \in [k, N - 1]_d$$

where  $f_{cl}^k(t, x) = f(x, \phi_k(t, x))$ .

- Instead of propagating uncertainty using the linearized model of the nonlinear closed-loop dynamics, we will use the *unscented transform*\*
- The unscented transform predicts future mean and covariance by only propagating  $2n + 1$  points known as the *sigma points*:

$$\begin{aligned} \sigma_k^0 &= \mu_k, \\ \sigma_k^i &= \mu_k + \sqrt{(n + \lambda)\Sigma_k^{1/2}} e_i, & \text{if } i \in [1, n]_d \\ \sigma_k^i &= \mu_k - \sqrt{(n + \lambda)\Sigma_k^{1/2}} e_i, & \text{if } i \in [n + 1, 2n]_d \end{aligned}$$

\* Wan, Eric A., and Rudolph Van Der Merwe. "The unscented Kalman filter for nonlinear estimation." Proc. of the IEEE 2000 Adaptive Systems for Signal Proc., Comm. and Control Symposium, 2000.



# Unscented transform for one-stage prediction of state mean and covariance

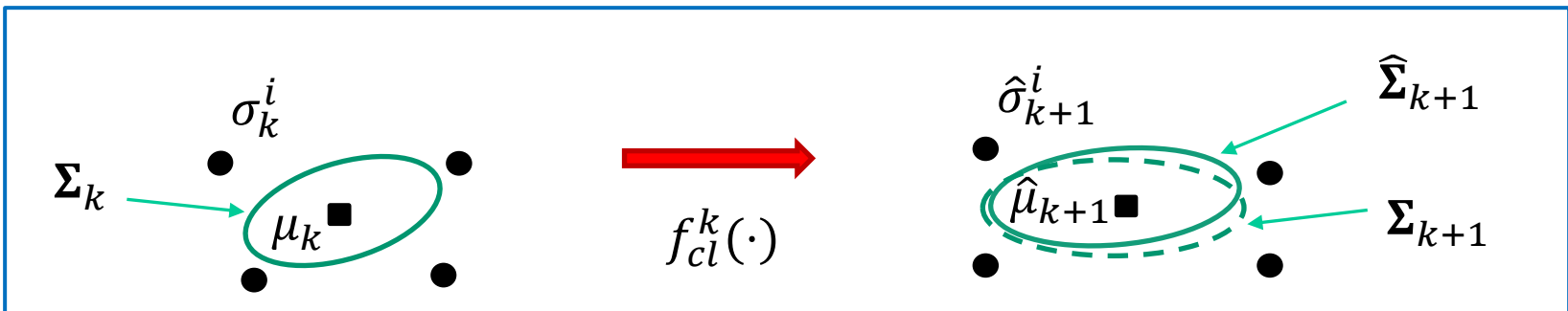
- Propagate the sigma points  $\sigma_k^i$  (stage  $t = k$ ) to obtain a new set of sigma points  $\hat{\sigma}_{k+1}^i$  (stage  $t = k + 1$ ):

$$\hat{\sigma}_{k+1}^i = f_{cl}^k(\sigma_k^i), \quad i \in [0, 2n]_d$$

- Compute approximations of the state mean and covariance at stage  $t = k + 1$  (one-stage predictions) by using the following equations:

$$\hat{\mu}_{k+1} = \sum_{i=0}^{2n} \gamma_k^i \hat{\sigma}_{k+1}^i$$

$$\hat{\Sigma}_{k+1} = \sum_{i=0}^{2n} \delta_k^i (\hat{\sigma}_{k+1}^i - \hat{\mu}_{k+1})(\hat{\sigma}_{k+1}^i - \hat{\mu}_{k+1})' + W_k$$



# Greedy Covariance Steering Algorithm

The proposed greedy (nonlinear) covariance steering algorithm\* consists of the following main steps (repeated until  $t = N - 1$ )

**Step 1:** Compute linearization  $(\mathbf{A}_k, \mathbf{B}_k, r_k)$  at stage  $t = k$  based on known approximations  $\hat{\mu}_k$  and  $\hat{\Sigma}_k$

**Step 2:** Compute control policy  $\pi_k = \{\phi_k(k, z), \dots, \phi_k(N - 1, z)\}$  that solves the  $k$ -th linearized covariance steering problem from  $(\hat{\mu}_k, \hat{\Sigma}_k)$  at  $t = k$  to  $(\mu_d, \Sigma_d)$  at  $t = N$

**Step 3:** Add  $\phi_k(k, z)$  (first control law of local control policy  $\pi_k$ ) to the “global” control policy  $\varpi$

**Step 4:** Compute one-stage predictions  $\hat{\mu}_{k+1}$  and  $\hat{\Sigma}_{k+1}$  via unscented transform

**Output:** Global control policy for nonlinear covariance steering

$$\varpi = \{\varphi_0(x), \varphi_1(x), \dots, \varphi_{N-1}(x), \}, \quad \varphi_k(x) = \phi_k(k, x)$$

\*Bakolas, Efstathios, and Alexandros Tsolovikos. "Greedy finite-horizon covariance steering for discrete-time stochastic nonlinear systems based on the unscented transform," *American Control Conference (ACC)*, 2020.



# Numerical Simulations (SDP Formulation)

- Consider a discrete-time nonlinear stochastic system
 
$$x(t+1) = x_1(t) + \tau x_2(t)$$

$$x_2(t+1) = x_2(t) - \tau(\delta x_1(t) + \zeta x_1(t)^3 + \gamma x_2(t)) + \tau u + \sqrt{\tau} w(t)$$
- Boundary conditions:  $x_0 \sim N(\mu_0, \Sigma_0)$ ,  $\mu_0 = 0$ ,  $\Sigma_0 = \text{diag}(0.4^2, 0.3^2)$ ,  $x_d \sim N(\mu_d, \Sigma_d)$ ,  $\mu_d = 0$  and  $\Sigma_d = \text{diag}(0.8^2, 0.6^2)$
- Parameters for numerical simulations:  $\zeta = \gamma = \alpha = 0.05$ ,  $\delta = -1$ ,  $\beta = 2$ , sampling period:  $\tau = 0.1$ , system parameters:  $N = 100$

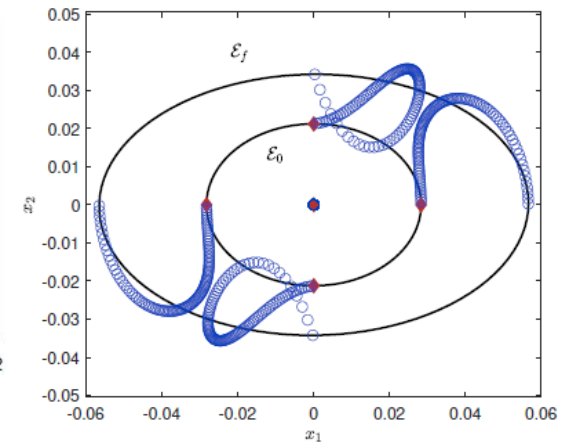
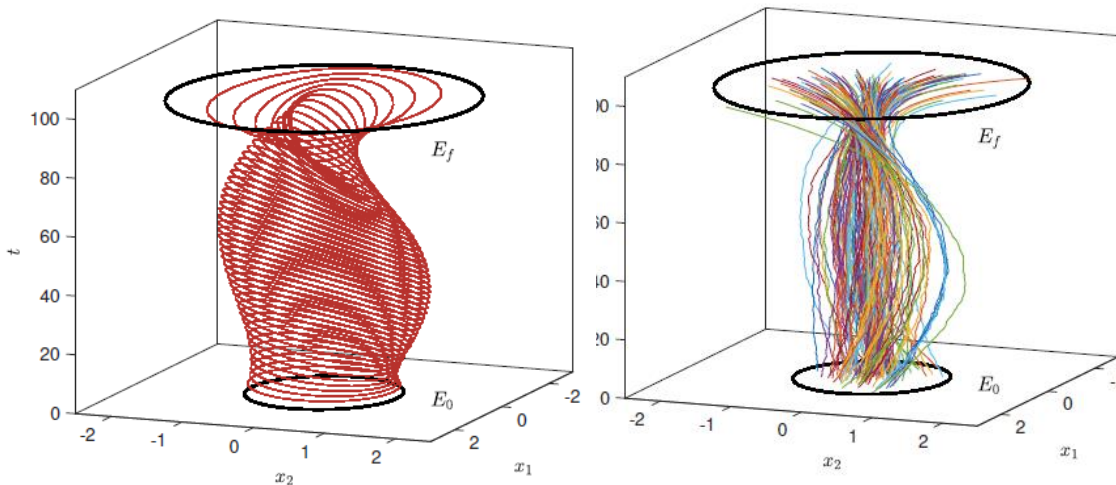


Figure 1. Three-dimensional illustration of the time-evolution of sample trajectories and state covariance of a nonlinear stochastic system

Figure 2. Time-evolution of the sigma points of the closed-loop system

# Numerical Simulations (DCP Formulation)

- Consider the discrete-time nonlinear stochastic model (unicycle car):

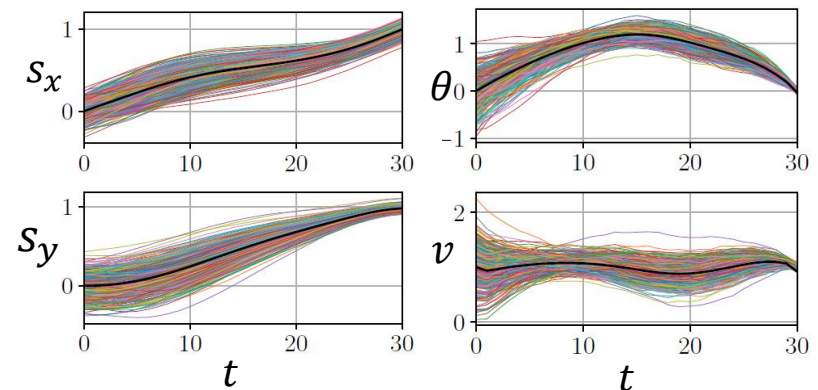
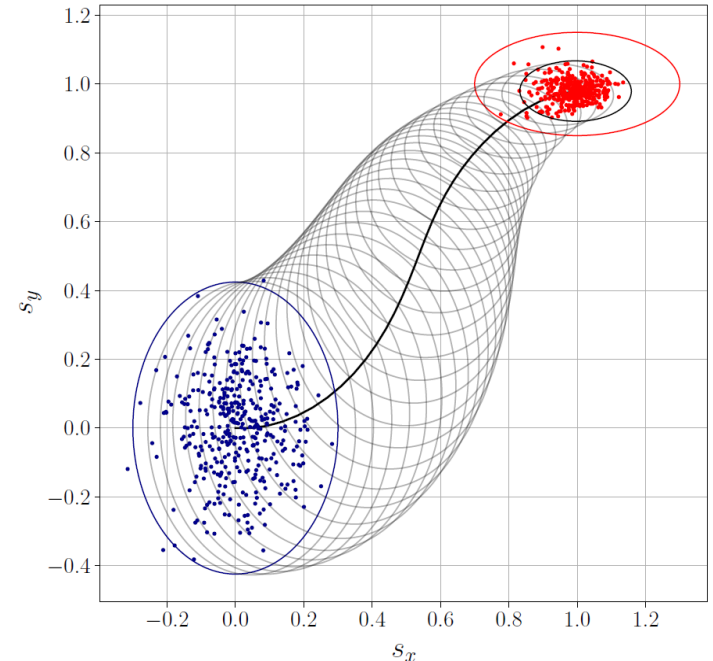
$$s_x(t+1) = s_x(t) + v(t)\tau\cos\theta(t) + \epsilon_x(t)$$

$$s_y(t+1) = s_y(t) + v(t)\tau\sin\theta(t) + \epsilon_y(t)$$

$$\theta(t+1) = \theta(t) + u_\theta(t)v(t)\tau + \epsilon_\theta(t)$$

$$v(t+1) = v(t) + u_v(t)\tau + \epsilon_v(t)$$

- $(s_x, s_y)$ : position vector,  $\theta$ : heading angle,  $v$ : speed
- Goal:** Shrink the uncertainty in the coordinate  $s_y$ , the heading angle  $\theta$ , and the speed  $v$ , while retaining the uncertainty in  $s_x$ .
- Terminal covariance (red) is close to the desired one even though we only considered soft terminal constraints (terminal cost  $\mathcal{W}^2(\rho_d, \rho_N)$ )





# Model-free Covariance Steering Based on Variational Gaussian Process Regression



# Covariance Steering Based on Data-Driven Predictive Models

- The previous tools for covariance steering assumed knowledge of a state space model of the uncertain system
- Such models may not be available a priori in many applications (e.g., the system dynamics may change during system operation)
- Proposed approach: Compute data-driven prediction models by learning the system dynamics from available data (experiment, simulations) and use these models for control design purposes
- In particular, use sparse variational Gaussian Process regression tools to capture the effects of model uncertainty and process noise (small computational / inference cost)



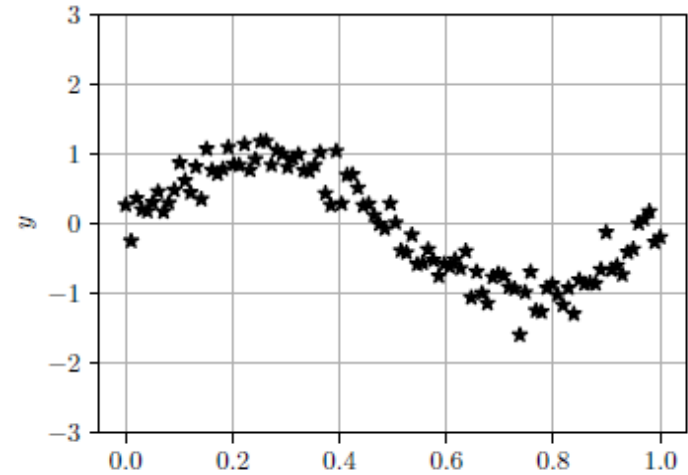
# Brief Introduction to Gaussian Processes for (non-parametric) Modeling

- Unknown function  $f: \mathbb{R}^n \rightarrow \mathbb{R}$  with noisy observations at known locations  $x_i$ :

$$y_i = f(x) + \epsilon_i$$

- Observation likelihood:

$$p(y_i | f(x_i)) = \mathcal{N}(y_i | f(x_i), \sigma_\epsilon^2)$$



- Data:  $N$  observations at  $N$  locations:  $y = \in \mathbb{R}^N$ ,  $\mathbf{X} = [x_1, \dots, x_N]^T \in \mathbb{R}^{N \times n} [y_1, \dots, y_N]^T$

## Why Gaussian Processes?

- Flexibility (non-parametric approach: distributions over functions)
- Provide uncertainty estimates
- Degrade gracefully (they know what they don't know)

## Basic Concepts and Steps of GP regression

- Assume that  $f$  belongs to a family of functions with a Gaussian Prior:

$$f(x) \sim N(f(x) | m(x), k(x, x))$$

- Prior over the vector  $\mathbf{f} = [f(x_1), \dots, f(x_N)]^T$ :

$$p(\mathbf{f}; \mathbf{X}) = \mathcal{N}(\mathbf{f} | m(\mathbf{X}), k(\mathbf{X}, \mathbf{X}))$$

- Joint density of  $\mathbf{y}$  and  $\mathbf{f}$ :  $p(\mathbf{y}, \mathbf{f}; \mathbf{X}) = p(\mathbf{y} | \mathbf{f}, \mathbf{X}) p(\mathbf{f}; \mathbf{X})$
- Marginalize out  $\mathbf{f}$  to obtain *marginal likelihood*:

$$p(\mathbf{y}; \mathbf{X}) = \int p(\mathbf{y} | \mathbf{f}; \mathbf{X}) p(\mathbf{f}; \mathbf{X}) d\mathbf{f} = \mathcal{N}(\mathbf{y} | m(\mathbf{X}), k(\mathbf{X}, \mathbf{X}) + \sigma_\epsilon^2 I)$$

- Optimize *hyperparameters*  $\Theta_\star = \{\theta_m, \theta_k, \sigma_\epsilon\}$ :

$$\Theta_\star = \underset{\Theta}{\operatorname{argmin}}(-\log(p(\mathbf{y}; \mathbf{X})))$$

where  $\theta_m, \theta_k$  are the mean and kernel hyperparameters



# Basic Concepts and Steps of GP regression

- *Prediction* (predict  $y_*$  on a test location  $x_*$ ):

$$p(y_*; x_*, \mathbf{y}, \mathbf{X}) = \int p(\mathbf{y}_*, \mathbf{y}; x_*, \mathbf{X}) d\mathbf{y} = N(\mathbf{y}_* | \mu_*, \sigma_*)$$

$$- \mu_* = m(x_*) + k(x_*, \mathbf{X})[k(\mathbf{X}, \mathbf{X}) + \sigma_\epsilon^2 I]^{-1}(\mathbf{y} - m(\mathbf{X}))$$

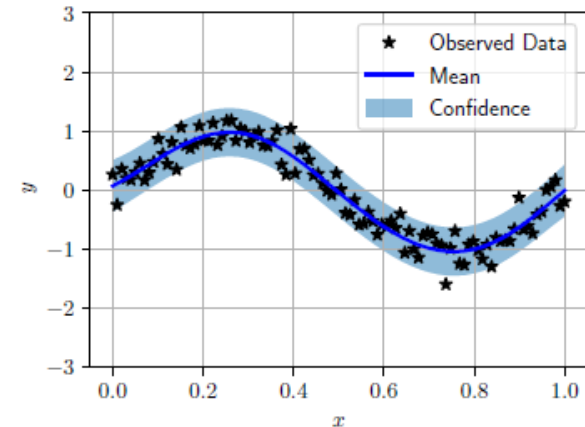
$$- \sigma_* = k(x_*, x_*) - k(x_*, \mathbf{X})[k(\mathbf{X}, \mathbf{X}) + \sigma_\epsilon^2 I]^{-1}k(\mathbf{X}, x_*)$$

- *Inference*: invert  $N \times N$  matrix (scales with cube of data size  $N$ ). Does not scale to more than a few thousand data points

- Example:

$$m(x) = c,$$

$$k(x, x') = \sigma_f^2 \exp\left(-\frac{1}{2}(x - x')^T \mathbf{L}^{-1}(x - x')\right)$$



# Scaling GP regression to Big Data: Sparse Variational GP Regression

- Sparse approximation of GPs (goal: reduce cost of inference)
- Introduce  $M$  inducing locations/values ( $M < N$ ):

$$\mathbf{Z} = [z_1 \dots z_M]^T, \quad \mathbf{u} = [f(z_1), \dots, f(z_M)]^T$$

- Joint density:  $p(\mathbf{y}, \mathbf{f}, \mathbf{u}) = p(\mathbf{y} | \mathbf{f}; \mathbf{X}) p(\mathbf{f} | \mathbf{u}; \mathbf{X}, \mathbf{Z}) p(\mathbf{u}; \mathbf{Z})$ 
  - $p(\mathbf{f} | \mathbf{u}; \mathbf{X}, \mathbf{Z}) = \mathcal{N}(\mathbf{f} | \tilde{\boldsymbol{\mu}}, \tilde{\boldsymbol{\Sigma}})$ , where
    - $[\tilde{\boldsymbol{\mu}}]_i = m(x_i) + k(x_i, \mathbf{Z})k(\mathbf{Z}, \mathbf{Z})^{-1}(u - m(\mathbf{Z}))$
    - $[\tilde{\boldsymbol{\Sigma}}]_{ij} = k(x_i, x_j) - k(x_i, \mathbf{Z})k(\mathbf{Z}, \mathbf{Z})^{-1}k(\mathbf{Z}, x_j)$
  - Gaussian prior on  $\mathbf{u}$ :  $p(\mathbf{u}; \mathbf{Z}) = \mathcal{N}(\mathbf{u} | m(\mathbf{Z}), k(\mathbf{Z}, \mathbf{Z}))$
- Variational posterior:  $q(\mathbf{f}, \mathbf{u}) = p(\mathbf{f} | \mathbf{u}; \mathbf{X}, \mathbf{Z})q(\mathbf{u})$   
 where  $q(\mathbf{u}) = \mathcal{N}(\mathbf{u} | m_u, \mathbf{S}_u)$  ( $m_u, \mathbf{S}_u$ : are parameters)

# Scaling GP regression to Big Data: Sparse Variational GP Regression

- Marginalize out  $\mathbf{u}$ :

$$q(\mathbf{f}|m_{\mathbf{u}}, \mathbf{S}_{\mathbf{u}}; \mathbf{X}, \mathbf{Z}) = \int p(\mathbf{f}|\mathbf{u}; \mathbf{X}, \mathbf{Z})q(\mathbf{u})d\mathbf{u} = \mathcal{N}(\mathbf{f}|\mu(\mathbf{X}), \Sigma(\mathbf{X}, \mathbf{X}))$$

where  $[\mu(\mathbf{X})]_i = \mu_f(x_i)$  and  $[\Sigma(\mathbf{X}, \mathbf{X})]_{i,j} = \Sigma_f(x_i, x_j)$ , with  $\mu_f$  and  $\Sigma_f$  defined as before\* (based on  $\mathbf{u}, \mathbf{Z}$ )

- Find optimal *variational parameters*  $\mathbf{Z}, m_{\mathbf{u}}$  and  $\mathbf{S}_{\mathbf{u}}$  and *hyperparameters* that maximize the following lower bound:

$$\log p(\mathbf{y} | \mathbf{X}) \geq \mathbb{E}_{q(\mathbf{f}, \mathbf{u})}[\log(p(\mathbf{y}, \mathbf{f}, \mathbf{u})/q(\mathbf{f}, \mathbf{u}))] = \mathcal{L}$$

$$\text{where } \mathcal{L} = \sum \mathbb{E}_{q(f_i|m, \mathbf{S}, x_i, \mathbf{Z})}[\log p(y_i|f_i) - \text{KL}(q(\mathbf{u})||p(\mathbf{u}))]$$

- Predict  $y_*$  on a test location  $x_*$ :

$$p(y_*; x_*, m, \mathbf{S}, \mathbf{Z}) = \mathcal{N}(y_*|\mu_f(x_*), \Sigma_f(x_*, x_*) + \sigma_{\epsilon}^2)$$

\*Tsolovikos, Alexandros, and Efstathios Bakolas. "Nonlinear Covariance Steering using Variational Gaussian Process Predictive Models." *arXiv preprint arXiv:2010.00778* (2020).

# System Identification using SVGPs

- **Given:**

1. Given a black box simulator corresponding to the originally unknown discrete-time nonlinear stochastic system:

$$z(t + 1) = g(z(t), u(t)) + \epsilon(t)$$

2. Observation data at known locations  $x_i = [z(t_i); u(t_i)]$ :

$$y_i = g(z(t_i), u(t_i)) + \epsilon(t_i),$$

- **Objective:**

Use data  $D = \{y_i, z_i, u_i\}_{i=1}^N$  to learn a SVGP-based prediction model for the system dynamics:

$$z(t + 1) = G(z(t), u(t)) + w(t)$$

where  $G(z(t), u(t)) := \mu_f([z(t); u(t)])$  and

$$w(t) \sim N(w_t | 0, \Sigma_f([z(t); u(t)], [z(t), u(t)]) + \sigma_\epsilon^2)$$



# Modifications to Model-Based Greedy Nonlinear Covariance Steering Algorithm

- Successive linearization of the SVGP-based predictive model:

$$z(t + 1) = A_{GP}z(t) + B_{GP}u(t) + d_{GP}$$

$$A_{GP} = \frac{\partial}{\partial z} G(\mu_z(t), \mu_u(t)), \quad B_{GP} = \frac{\partial}{\partial u} G(\mu_z(t), \mu_u(t))$$

$$d_{GP} = -A_{GP}\mu_z(t) - B_{GP}\mu_u(t) + G(\mu_z(t), \mu_u(t))$$

- The unscented transform will also be adjusted\*. In particular,
  - GP-based predictive model will be used for sigma points propagation:

$$\hat{\sigma}_{k+1}^i = G_{cl}^k(\sigma_k^i), \quad G_{cl}^k(t, z) := G(z, \phi_k(t, z))$$

- The noise covariance used in the computation of the next stage covariance will also be adjusted appropriately

$$\hat{\Sigma}_{k+1} = \sum_{i=0}^{2n} \delta_k^i (\hat{\sigma}_{k+1}^i - \hat{\mu}_{k+1})(\hat{\sigma}_{k+1}^i - \hat{\mu}_{k+1})' + W_k$$

where  $W_k = \Sigma_f([z(t); u(t)], [z(t); u(t)]) + \sigma_\epsilon^2 I$

\* Ko, Jonathan, et al. "GP-UKF: Unscented Kalman filters with Gaussian process prediction and observation models." 2007

IEEE/RSJ International Conference on Intelligent Robots and Systems, 2007.

# Simulation Results

- Consider again the system:

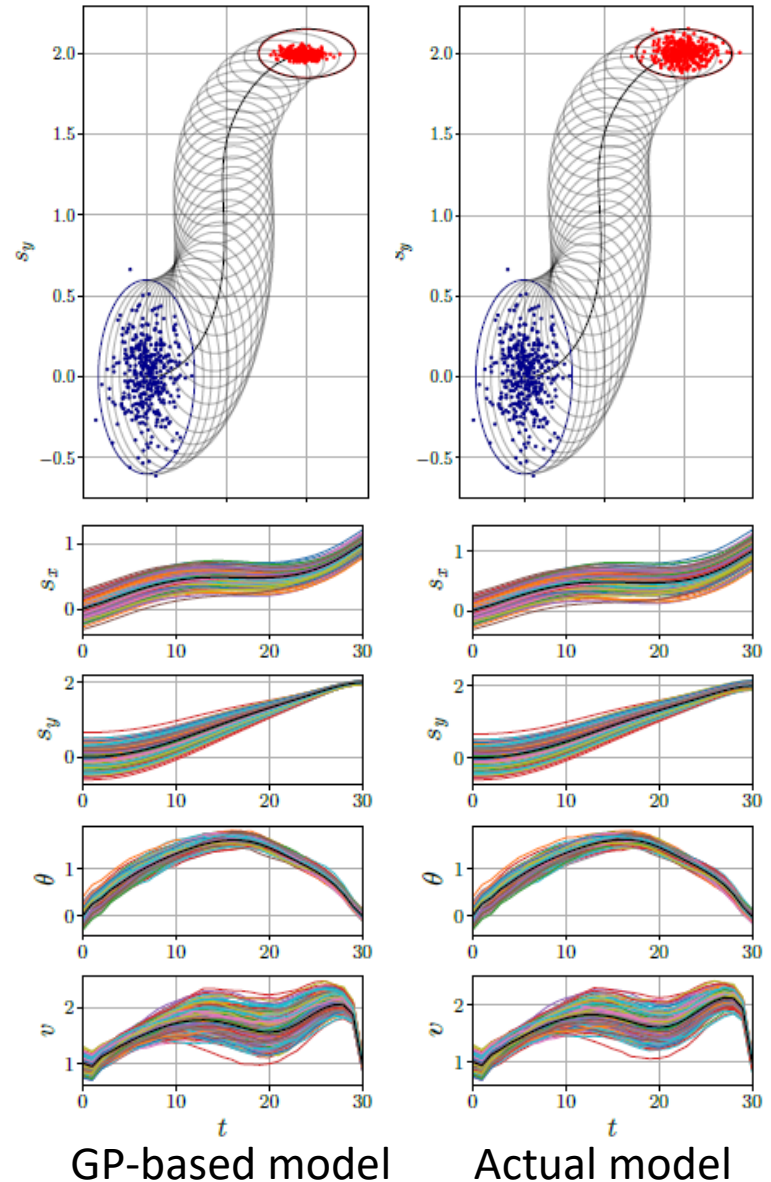
$$s_x(t + 1) = s_x(t) + v(t)\tau\cos\theta(t) + \epsilon_x(t)$$

$$s_y(t + 1) = s_y(t) + v(t)\tau\sin\theta(t) + \epsilon_y(t)$$

$$\theta(t + 1) = \theta(t) + u_\theta(t)v(t)\tau + \epsilon_\theta(t)$$

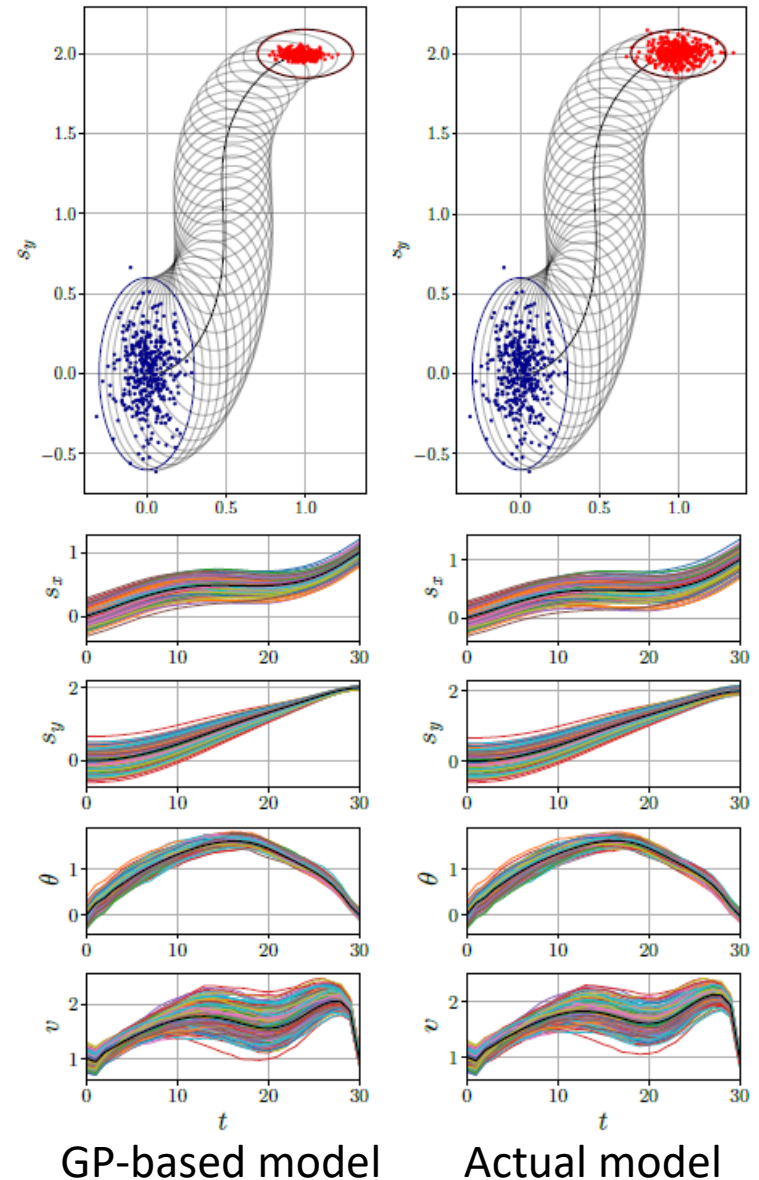
$$v(t + 1) = v(t) + u_v(t)\tau + \epsilon_v(t)$$

- $(s_x, s_y)$ : position vector,  $\theta$ : heading angle,  $v$ : speed
- Data from a black box simulator are used for training a GP-based prediction model (use squared exponential kernel)
- Goal:** Shrink the uncertainty in the coordinate  $s_y$ , the heading angle  $\theta$ , and the speed  $v$ , while retaining the uncertainty in  $s_x$ .



# Remarks on numerical simulations

- Left figures: results based on the SVGP-prediction models; Right figures: results based on model-based covariance steering (SDP formulation)
- The uncertainty predicted by the SVGP model is very close to the uncertainty predicted by the model-based approach.
- The actual terminal distribution obtained using SVGP model (visualization based on red particles from 400 Monte Carlo realizations) is more concentrated near the mean than in the model-based approach (right figure).
- The covariance steering based on the SVGP model is more *cautious* than model-based covariance steering.



## Summary & Concluding Remarks

- We discussed ways to address covariance / distribution steering problems for discrete-time nonlinear stochastic systems using model-based and model-free (data-driven) approaches
  - successive linearization of system dynamics (based on either a given state-space model or a SVGP-based predictive model) along the ensuing mean (state and input) trajectories
  - the solution of a sequence of linearized covariance steering problems which are either associated with SDP (convex) programs or difference of convex functions programs
  - the unscented transform for the computation of one-stage predictions of the state mean and state covariance



## Opportunities for Future Research

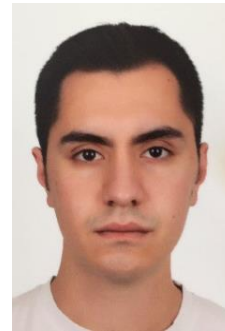
- Study the partial state information case for both the model-based and the model-free cases
- Improve performance of linearized covariance steering algorithms (consider different control policy parametrizations)
- Explore different linearization methods in order to reduce the frequency of linearization
- Explore better ways to improve scalability and computational efficiency, and make connections with nonlinear MPC methods (infinite-horizon case)
- Study the distribution steering problems in the class of multimodal distributions



- PhD students who contributed to this research



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